

Econophysics III: Financial Correlations and Portfolio Optimization

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Let's Face Chaos through Nonlinear Dynamics, Maribor 2011

Outline

Portfolio optimization is a key issue when investing money. It is applied science and everyday work for many physicists who join the financial industry.

- importance of **financial correlations**
- portfolio optimization with **Markowitz theory**
- problem of **noise dressing**
- cleaning methods **filtering** and **power mapping**
- application to **Swedish and US market data**
- rôle of **constraints** in portfolio optimization

Portfolio, Risk and Correlations

Putting together a Portfolio

Portfolio 1

ExxonMobil

British Petrol

Daimler

Toyota

ThyssenKrupp

Voestalpine

Portfolio 2

Sony

British Petrol

Daimler

Coca Cola

Novartis

Voestalpine

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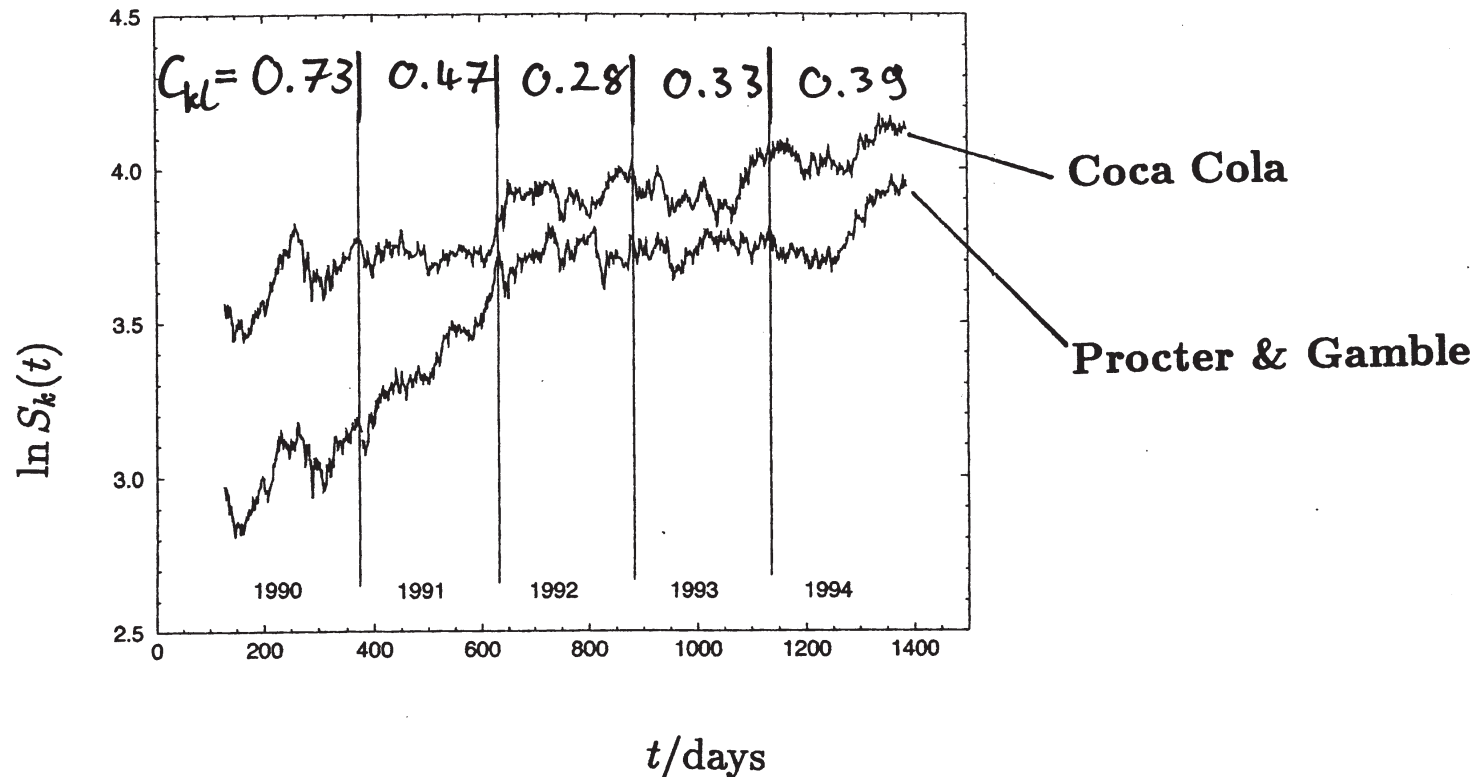
Novartis

Voestalpine

correlations \longrightarrow diversification lowers portfolio risk!

Correlations between Stocks

visual inspection for Coca Cola and Procter & Gamble



correlations change over time!

Portfolio and Risk Management

portfolio is linear combination of stocks, options and other financial instruments

$$V(t) = \sum_{k=1}^K w_k(t) S_k(t) + \sum_{l=1}^L w_{Cl}(t) G_{Cl}(S_l, t) + \sum_{m=1}^M w_{Pm}(t) G_{Pm}(S_m, t) + \dots$$

with **time-dependent** weights!

portfolio or fund manager has to maximize return

- high return requires high risk: **speculation**
- low risk possible with **hedging and diversification**

find optimum for risk and return according to investors' wishes

→ **risk management**

Risk of a Portfolio

general: $V(t) = \sum_{k=1}^K w_k(t) F_k(t)$ with risk elements $F_k(t)$

define moments within a time T

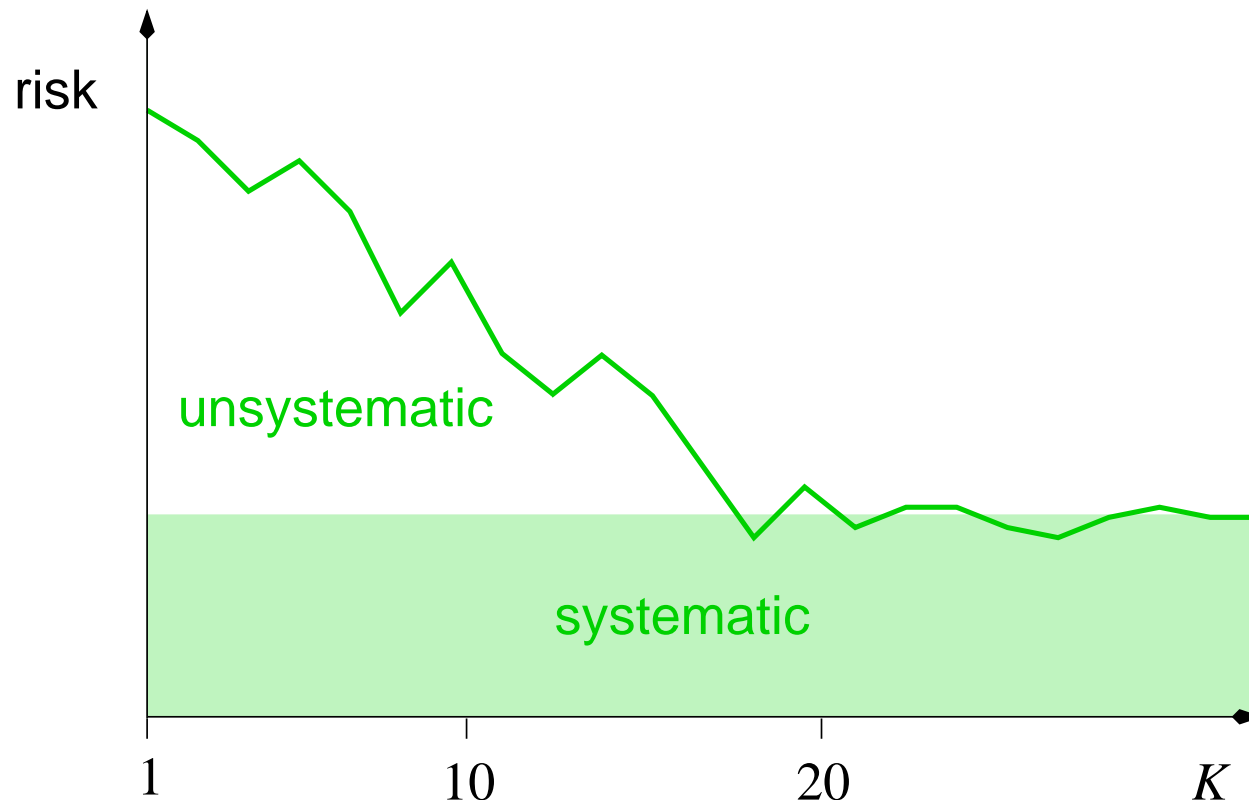
$$\langle V(t) \rangle = \frac{1}{T} \sum_{t=1}^T V(t) \quad \text{and} \quad \langle V^2(t) \rangle = \frac{1}{T} \sum_{t=1}^T V^2(t)$$

risk is the variance of the portfolio $\langle V^2(t) \rangle - \langle V(t) \rangle^2$

often normalized $\frac{\langle V^2(t) \rangle - \langle V(t) \rangle^2}{\langle V(t) \rangle^2}$

Diversification — Empirically

systematic risk (market) and unsystematic risk (portfolio specific)



a wise choice of $K = 20$ stocks (or risk elements) turns out sufficient to eliminate unsystematic risk

Risk, Covariances and Correlations

if the weights $w_k(t)$ are time-independent within the time interval T under consideration, one has

$$\begin{aligned}\langle V^2(t) \rangle - \langle V(t) \rangle^2 &= \sum_{k,l} w_k w_l \langle F_k(t) F_l(t) \rangle - \left(\sum_{k=1}^K w_k \langle F_k(t) \rangle \right)^2 \\ &= \sum_{k,l} w_k w_l \left(\langle F_k(t) F_l(t) \rangle - \langle F_k(t) \rangle \langle F_l(t) \rangle \right) \\ &= \sum_{k,l} w_k w_l \underbrace{\left\langle \left(F_k(t) - \langle F_k(t) \rangle \right) \left(F_l(t) - \langle F_l(t) \rangle \right) \right\rangle}_{\text{covariance matrix element } \Sigma_{kl}}\end{aligned}$$

covariance matrix element Σ_{kl}

Measuring Financial Correlations

stock prices $S_k(t)$, $k = 1, \dots, K$ for K companies measured at times $t = 1, \dots, T$

returns $G_k(t) = \ln \frac{S_k(t + \Delta t)}{S_k(t)} \simeq \frac{dS_k(t)}{S_k(t)}$

volatilities $\sigma_k(T) = \sqrt{\langle G_k^2(t) \rangle - \langle G_k(t) \rangle^2}$

normalized time series $M_k(t) = \frac{G_k(t) - \langle G_k(t) \rangle}{\sigma_k}$

correlation $C_{kl}(T) = \langle M_k(t) M_l(t) \rangle = \frac{1}{T} \sum_{t=1}^T M_k(t) M_l(t)$

Financial Correlation and Covariance Matrices

correlation matrix $C = C(T)$ is $K \times K$ with elements $C_{kl}(T)$

$$C = C(T) = \frac{1}{T} M M^\dagger$$

covariance $\Sigma_{kl}(T) = \left\langle (G_k(t) - \langle G_k(t) \rangle) (G_l(t) - \langle G_l(t) \rangle) \right\rangle$

covariance and correlation are related by

$$\Sigma_{kl}(T) = \sigma_k(T) C_{kl}(T) \sigma_l(T), \quad \text{such that} \quad \Sigma = \sigma C \sigma,$$

where $\sigma = \text{diag}(\sigma_1, \dots, \sigma_K)$ measured volatilities

Markowitz Portfolio Optimization

Portfolio Risk and Return

portfolio $V(t) = \sum_{k=1}^K w_k S_k(t) = w \cdot S(t)$

w_k fraction of wealth invested, normalization

$$1 = \sum_{k=1}^K w_k = w \cdot e \quad \text{with } e = (1, \dots, 1)$$

desired return for portfolio $R = \sum_{k=1}^K w_k r_k = w \cdot r$

$r_k = dS_k/S_k$ expected return for stock k

risk $\Omega^2 = \sum_{k,l} w_k \Sigma_{kl} w_l = w^\dagger \Sigma w = w^\dagger \sigma C \sigma w$

Optimal Portfolio

find those fractions $w_k^{(\text{opt})}$ which yield the desired return R at the minimum risk Ω^2

Euler–Lagrange optimization problem

$$L = \frac{1}{2} w^\dagger \Sigma w - \alpha (w \cdot r - R) - \beta (w \cdot e - 1)$$

α, β Lagrange multipliers, $\Sigma = \sigma C \sigma$ covariance

$$\longrightarrow \quad \frac{\partial L}{\partial w} = 0, \quad \frac{\partial L}{\partial \alpha} = 0, \quad \frac{\partial L}{\partial \beta} = 0$$

system of $K + 2$ equations

No Constraints: Closed Form Solutions

$K + 2$ coupled equations read explicitly

$$0 = \Sigma w^{(\text{opt})} - \alpha r - \beta e$$

$$0 = r \cdot w^{(\text{opt})} - R$$

$$0 = e \cdot w^{(\text{opt})} - 1$$

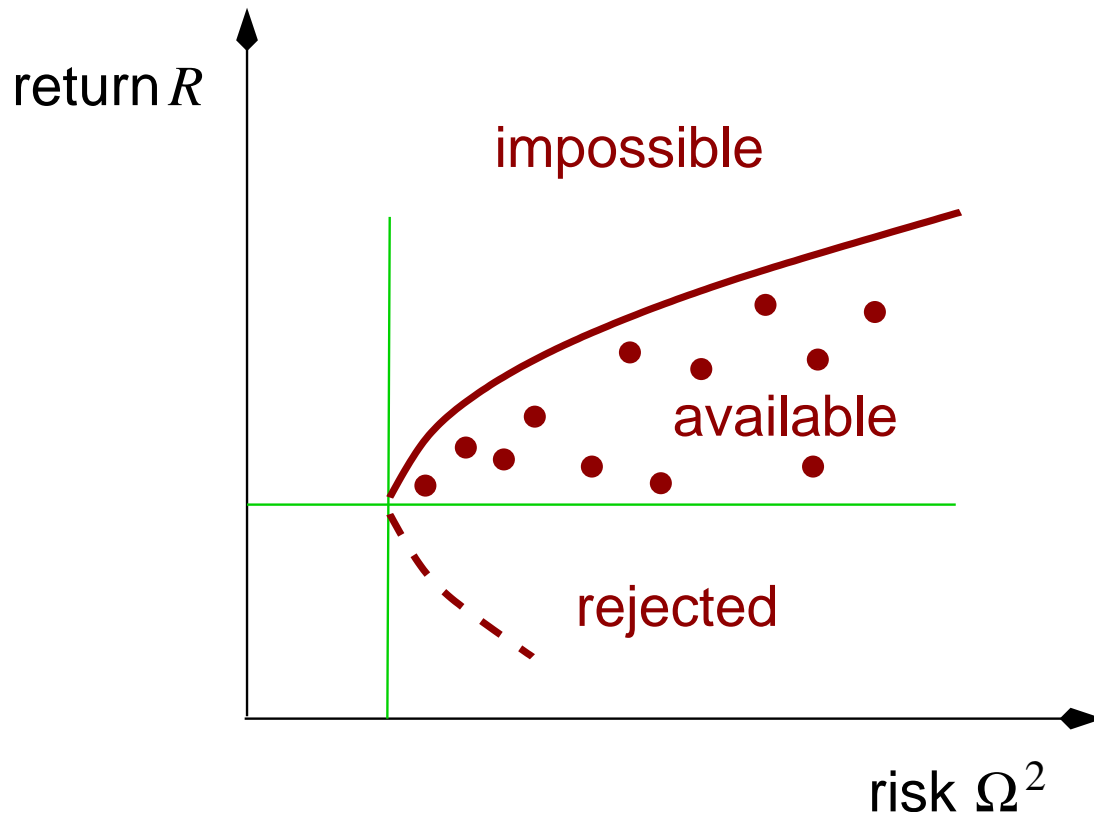
if they exist, solutions are given by

$$w^{(\text{opt})} = \frac{\Sigma^{-1}e}{e^\dagger \Sigma^{-1}e} + \frac{e^\dagger \Sigma^{-1}e R - r^\dagger \Sigma^{-1}e}{e^\dagger \Sigma^{-1}e r^\dagger \Sigma^{-1}r - (e^\dagger \Sigma^{-1}e)^2} \Sigma^{-1} \left(r - \frac{e^\dagger \Sigma^{-1}r}{e^\dagger \Sigma^{-1}e} e \right)$$

check that they minimize risk

risk Ω^2 is a quadratic function in desired return R

Efficient Frontier



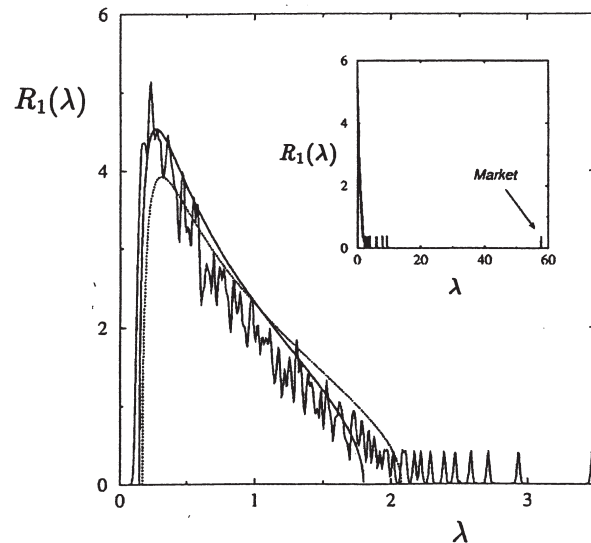
return R is a square root in the risk Ω^2

inclusion of further constraints possible, for example no short selling, $w_k \geq 0$ \longrightarrow efficient frontier changes

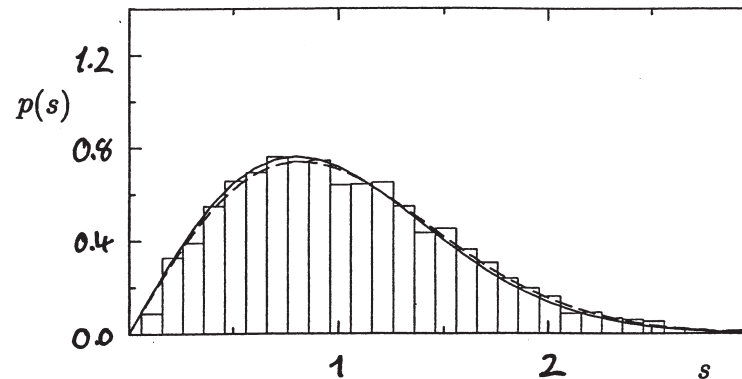
Noise Dressing of Financial Correlations

Empirical Results

correlation matrices of S&P500 and TAQ data sets



eigenvalue density



distribution of spacings
between the eigenvalues

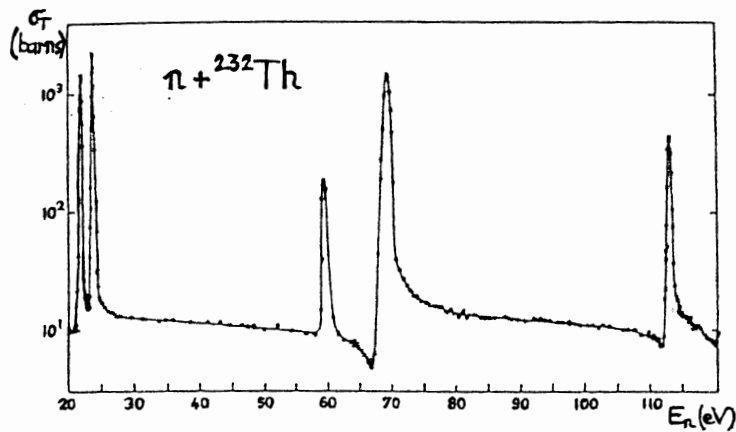
true correlations are noise dressed \longrightarrow DISASTER!

Laloux, Cizeau, Bouchaud, Potters, PRL 83 (1999) 1467

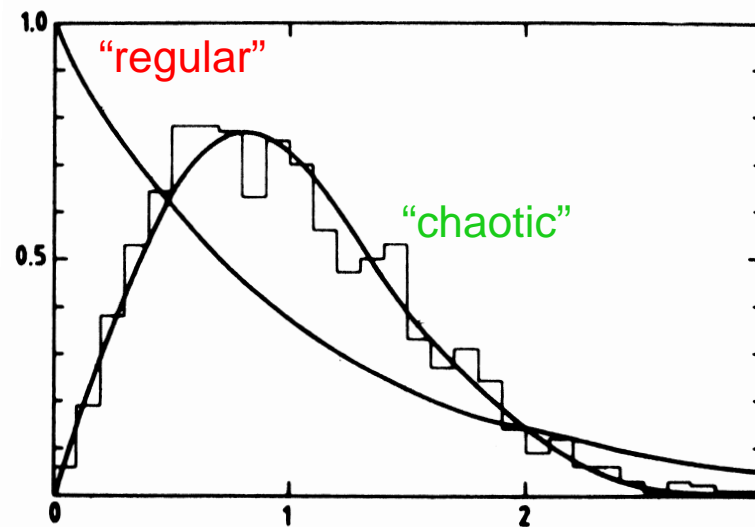
Plerou, Gopikrishnan, Rosenow, Amaral, Stanley, PRL 83 (1999) 1471

Quantum Chaos

result in statistical nuclear physics (Bohigas, Haq, Pandey)



resonances



spacing distribution

universal in a huge variety of systems: nuclei, atoms, molecules, disordered systems, lattice gauge quantum chromodynamics, elasticity, electrodynamics

→ quantum "chaos" → random matrix theory

Major Reason for the Noise

$$C_{kl} = \frac{1}{T} \sum_{t=1}^T M_k(t) M_l(t) \quad \text{we look at} \quad z_{kl} = \frac{1}{T} \sum_{t=1}^T a_k(t) a_l(t)$$

with uncorrelated standard normal time series $a_k(t)$

z_{kl} to leading order Gaussian distributed with variance T

$$z_{kl} = \delta_{kl} + \sqrt{\frac{1 + \delta_{kl}}{T}} \alpha_{kl} \quad \text{with standard normal } \alpha_{kl}$$

→ noise dressing $C = C_{\text{true}} + C_{\text{random}}$ for finite T

it so happens that C_{random} is equivalent to a random matrix in the chiral orthogonal ensemble

Chiral Random Matrices

Dirac operator in relativistic quantum mechanics, M is $K \times T$

$$D = \begin{bmatrix} 0 & M/\sqrt{T} \\ M^\dagger/\sqrt{T} & 0 \end{bmatrix}$$

chiral symmetry implies off-block diagonal form

eigenvalue spectrum follows from

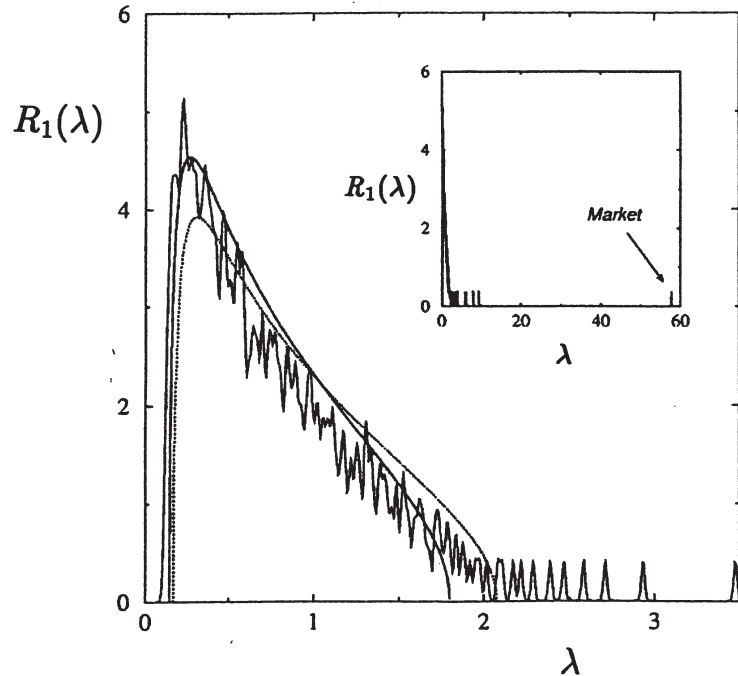
$$0 = \det(\lambda 1_{K+T} - D) = \lambda^{T-K} \det(\lambda^2 1_K - MM^\dagger/T)$$

where $C = MM^\dagger/T$ has the form of the correlation matrix

if entries of M are Gaussian random numbers, eigenvalue density

$$R_1(\lambda) = \frac{1}{2\pi\lambda} \sqrt{(\lambda_{\max} - \lambda)(\lambda - \lambda_{\min})}$$

Correlation Matrix is Largely Random



random matrix behavior

is here a DISASTER

serious doubts about practical usefulness of correlation matrices

... but: what is the meaning of the large eigenvalues ?

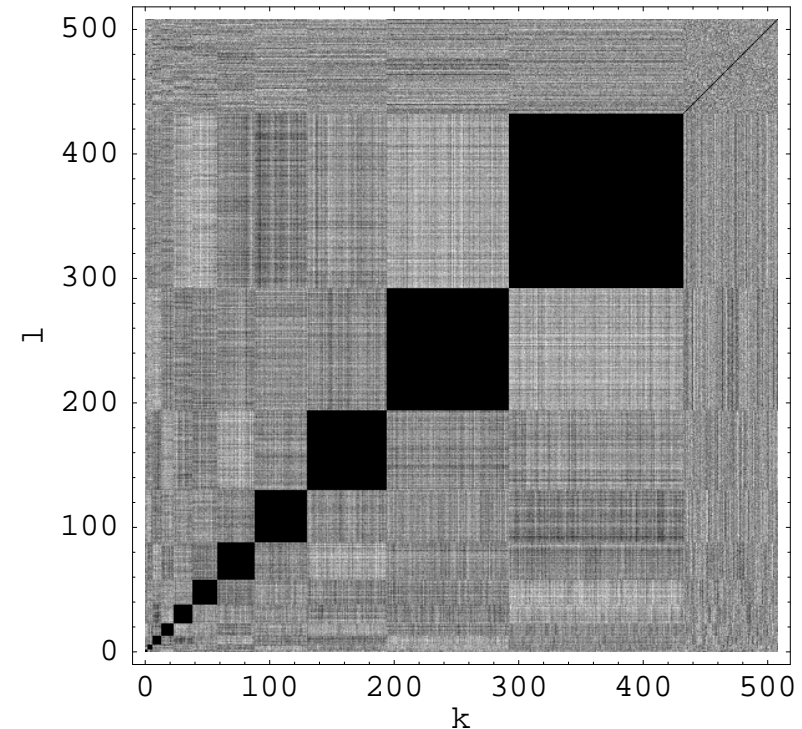
A Model Correlation Matrix

one-factor model (called Noh's model in physics)

$$M_k(t) = \frac{\sqrt{p_{b(k)}}\eta_{b(k)}(t)}{\sqrt{1 + p_{b(k)}}} + \frac{\varepsilon_k(t)}{\sqrt{1 + p_{b(k)}}}$$

branch plus idiosyncratic

$\eta_{b(k)}(t)$ and $\varepsilon_k(t)$ are standard normal, uncorrelated time series



Explanation of the Large Eigenvalues

$\kappa_b \times \kappa_b$ matrix $\begin{bmatrix} 1 & \dots & 1 \\ \vdots & & \vdots \\ 1 & \dots & 1 \end{bmatrix} = ee^\dagger$ with $e = \begin{bmatrix} 1 \\ \vdots \\ 1 \end{bmatrix}$

for $T \rightarrow \infty$, one block has the form $\frac{1}{1+p_b} (p_b ee^\dagger + 1_{\kappa_b})$

e itself is an eigenvector! — it yields large eigenvalue $\frac{1+p_b\kappa_b}{1+p_b}$

in addition, there are $\kappa_b - 1$ eigenvalues $\frac{1}{1+p_b}$

Noise Reduction and Cleaning

Noise Reduction by Filtering

diagonalize $K \times K$ correlation matrix $C = U^{-1}\Lambda U$

remove noisy eigenvalues $\Lambda = \text{diag}(\lambda_1, \dots, \lambda_c, \lambda_{c+1}, \dots, \lambda_K)$



keep branch eigenvalues $\Lambda^{(\text{filtered})} = \text{diag}(0, \dots, 0, \lambda_{c+1}, \dots, \lambda_K)$

obtain filtered $K \times K$ correlation matrix $C^{(\text{filtered})} = U^{-1}\Lambda^{(\text{filtered})}U$

restore normalization $C_{kk}^{(\text{filtered})} = 1$

Bouchaud, Potters, Theory of Financial Risk (2000)

Plerou, Gopikrishnan, Rosenow, Amaral, Guhr, Stanley, PRE 65 (2002) 066126

Parameter and Input Free Alternative?

What if “large” eigenvalue of a smaller branch lies in the bulk?

Also: For smaller correlation matrices, cut-off eigenvalue λ_c not so obvious.

There are many more noise reduction methods.

It seems that all these methods involve **parameters** to be chosen or other **input**.

Introduce the **power mapping** as an example for a new method.
It needs little input.

The method exploits the **chiral structure** and the **normalization**.

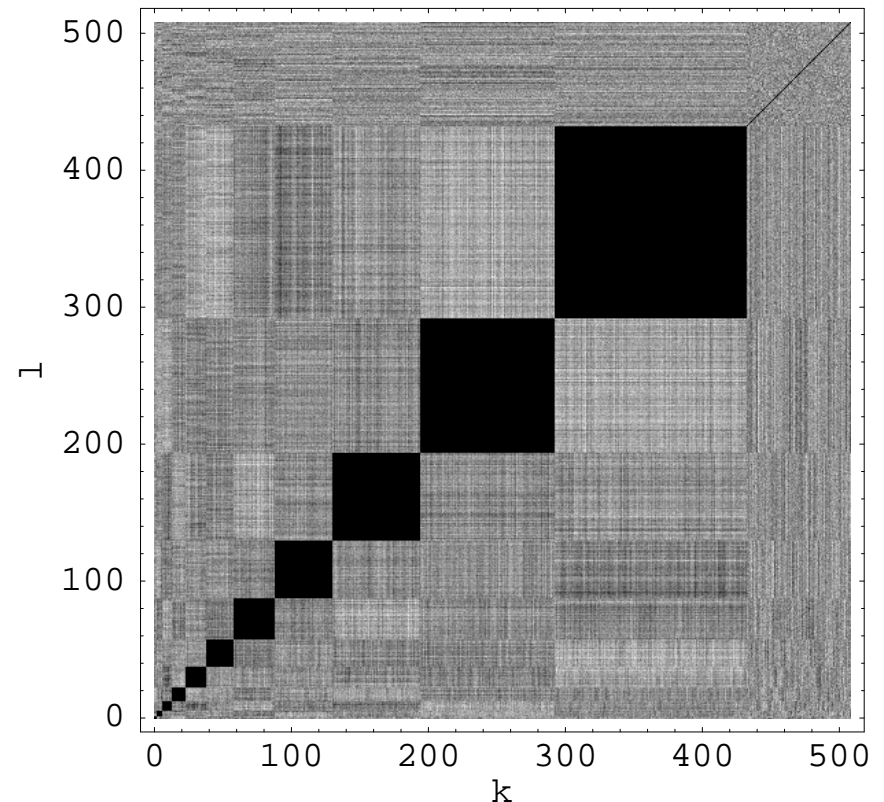
Illustration Using the Noh Model Correlation Matrix

We look at a **synthetic** correlation matrix.

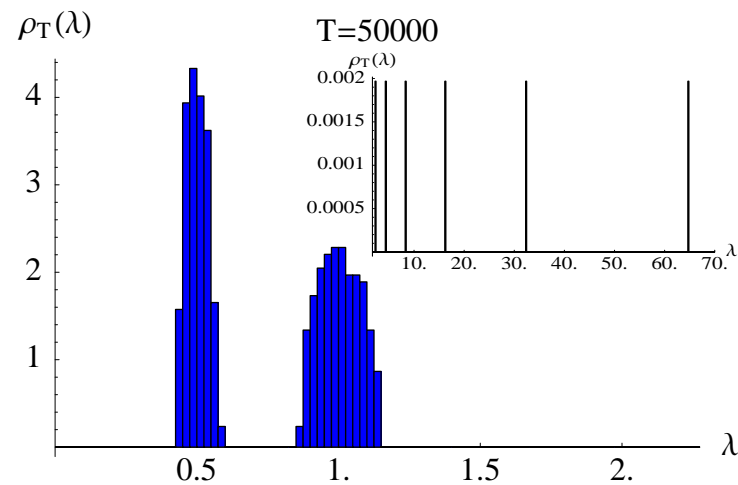
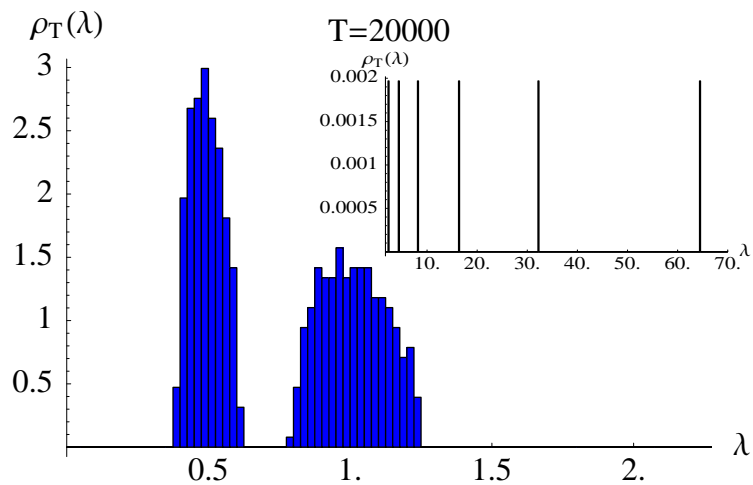
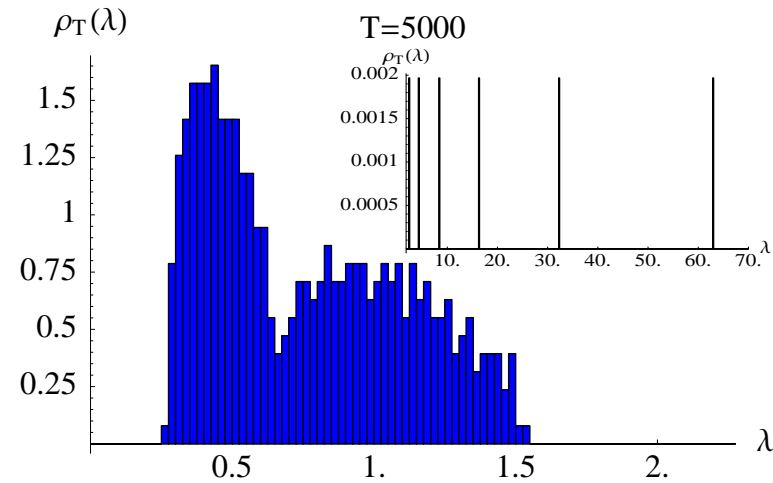
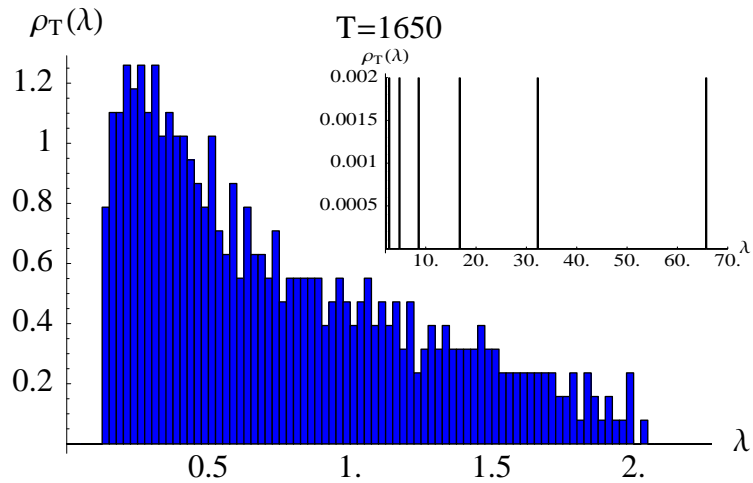
one-factor model (called Noh's model in physics)

$$M_k(t) = \frac{\sqrt{p_{b(k)}} \eta_{b(k)}(t)}{\sqrt{1 + p_{b(k)}}} + \frac{\varepsilon_k(t)}{\sqrt{1 + p_{b(k)}}}$$

branch plus idiosyncratic

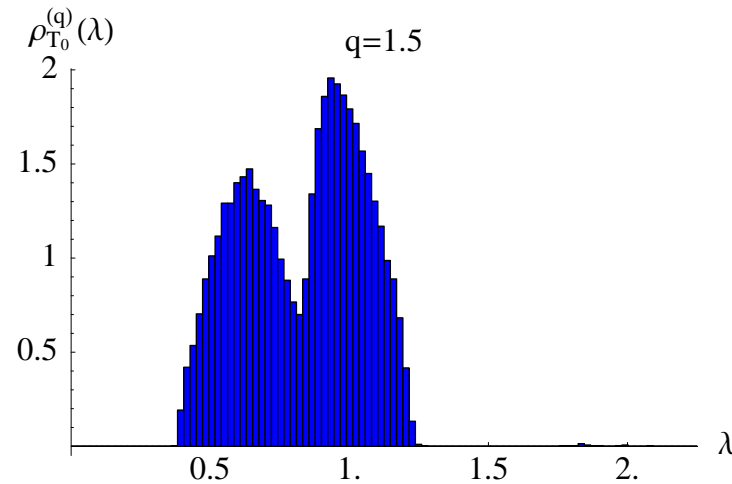
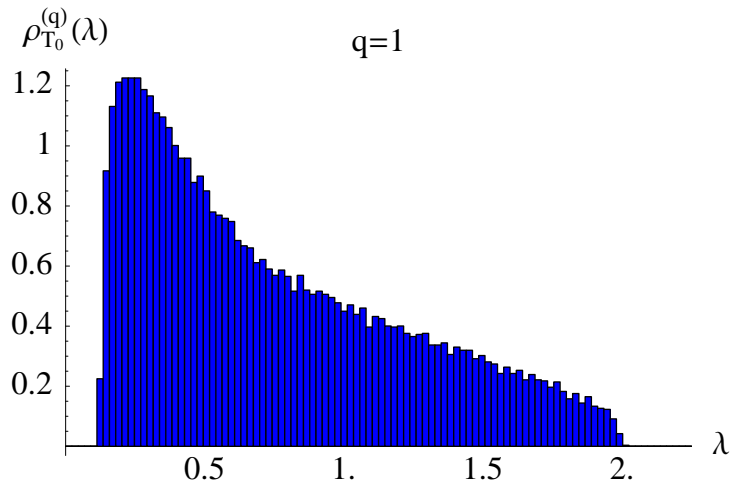


Spectral Densities and Length of the Time Series



→ correlations and noise separated

Power Mapping



$$C_{kl}(T) \quad \longrightarrow \quad \text{sign}(C_{kl}(T)) |C_{kl}(T)|^q$$

large eigenvalues (branches) only little affected

time series are effectively “prolonged” !

Heuristic Explanation

matrix element C_{kl} containing true correlation u and noise v/\sqrt{T}

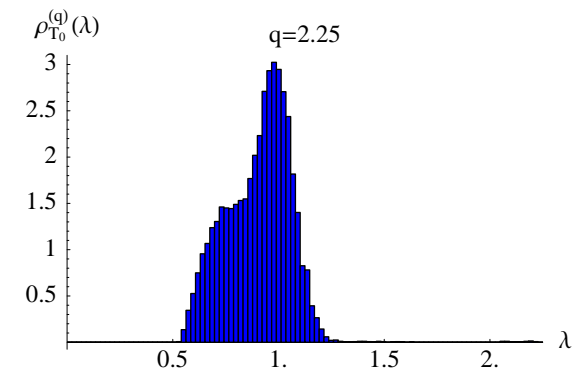
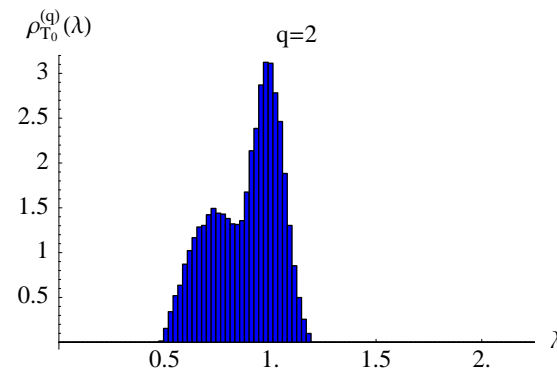
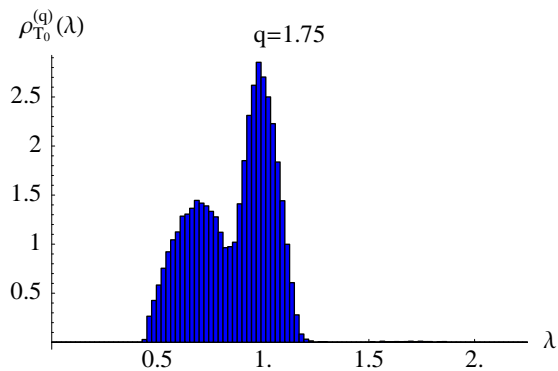
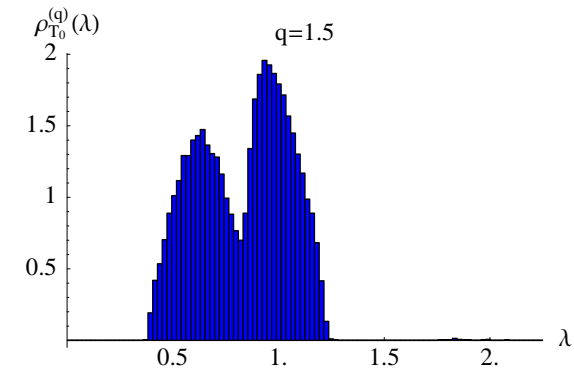
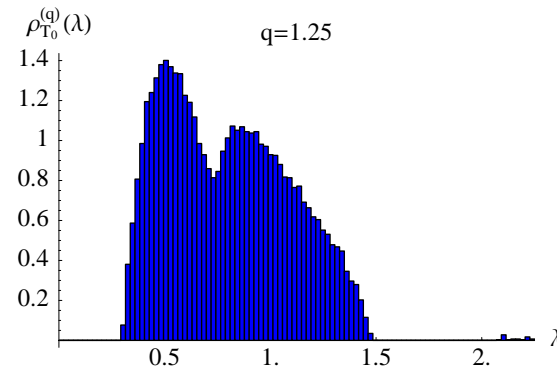
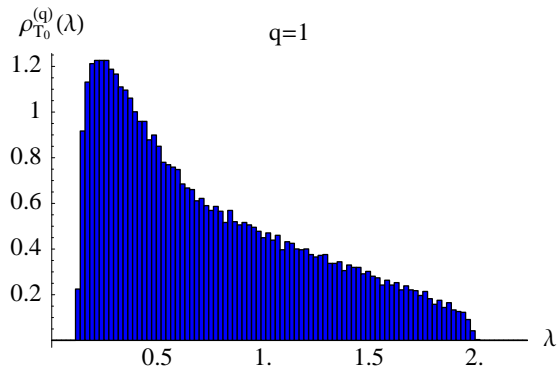
$$\left(u + \frac{v}{\sqrt{T}}\right)^q = u^q + q \frac{u^{q-1}v}{\sqrt{T}} + \mathcal{O}\left(\frac{1}{T}\right)$$

matrix element C_{kl} containing only noise v/\sqrt{T}

$$\left(\frac{v}{\sqrt{T}}\right)^q = \frac{v^q}{T^{q/2}}$$

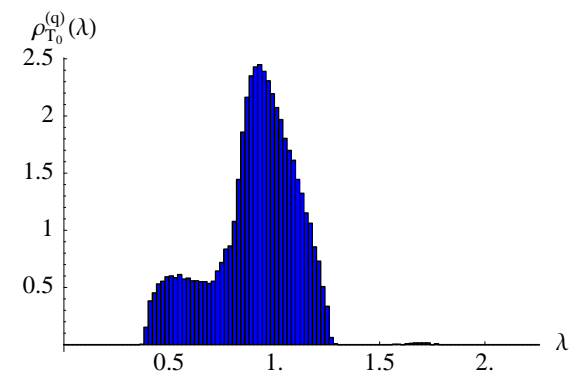
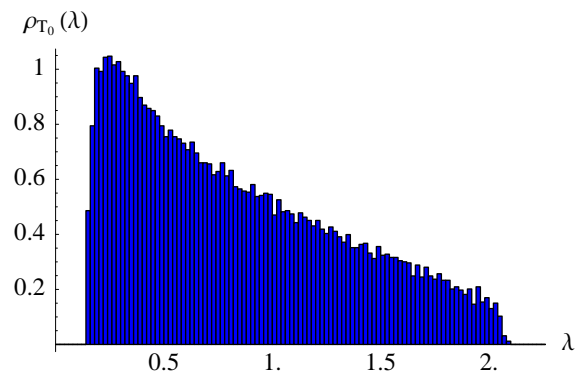
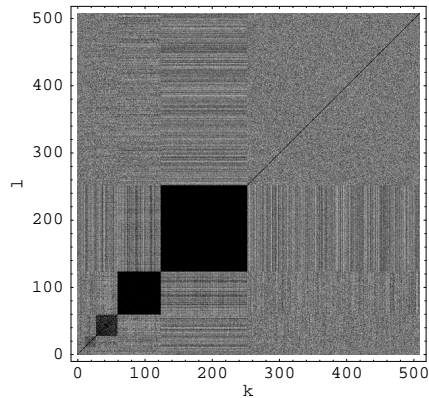
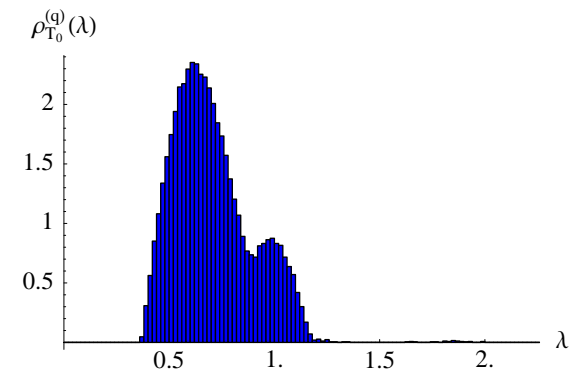
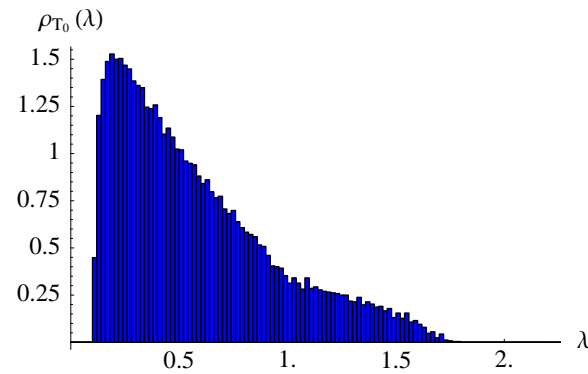
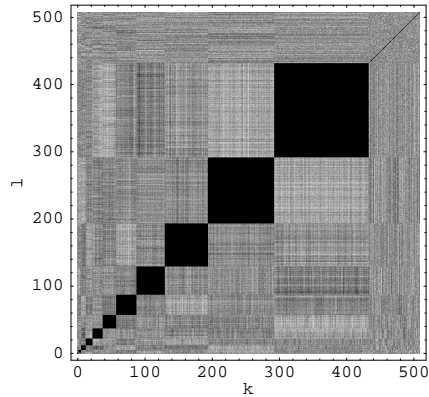
→ noise suppressed for $q > 1$

Self-determined Optimal Power



optimal power $q \approx 1.5$ is automatically determined by the very definition of the correlation matrix

Internal Correlation Structure



power mapping sensitive enough to clean the internal structure

Power Mapping is a New Shrinkage Method

shrinkage in mathematical statistics means removal of something which one does not want to be there (noise)

→ in practice: linear subtraction methods

→ shrinkage parameter (and other input) needed

power mapping is non-linear

it is parameter free and input free, because

- “chirality”

correlation matrix elements C_{kl} are scalar products

→ noise goes like $1/\sqrt{T}$ to leading order

- normalization

boundness $|C_{kl}| \leq 1 \quad \longrightarrow \quad |C_{kl}|^q \leq 1$

Sketch of Analytical Discussion

$\lambda_k(T)$ eigenvalues of $K \times K$ correlation matrix $C = C(T)$

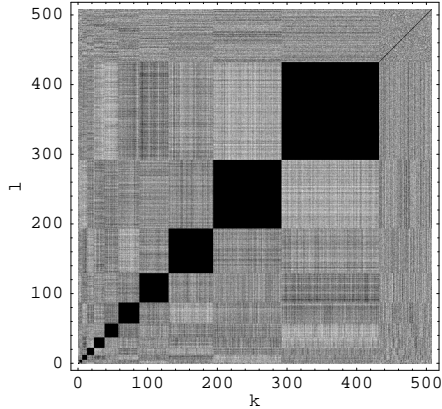
before power mapping $\lambda_k(T) = \lambda_k(\infty) + \frac{v_k}{\sqrt{T}} a_k + \mathcal{O}(1/T)$

$$\rho_T(\lambda) = \int_{-\infty}^{+\infty} d\lambda' \frac{1}{K} \sum_{k=1}^K G\left(\lambda - \lambda', \frac{v_k^2}{T}\right) \rho_\infty(\lambda') + \mathcal{O}(1/T)$$

thereafter $\lambda_k^{(q)}(T) = \lambda_k^{(q)}(\infty) + \frac{v_k^{(q)}}{\sqrt{T}} a_k^{(q)} + \frac{\tilde{v}_k^{(q)}}{T^{q/2}} \tilde{a}_k^{(q)} + \mathcal{O}(1/T)$

$$\rho_T^{(q)}(\lambda) = \int_{-\infty}^{+\infty} d\lambda' \frac{1}{K} \sum_{k=1}^K G\left(\lambda - \lambda', \frac{(\tilde{v}_k^{(q)})^2}{T^q}\right) \rho_T(\lambda') \Big|_{v_k}^{v_k^{(q)}} + \mathcal{O}(1/T)$$

Result for Power-Mapped Noh Model



$$M_k(t) = \frac{\sqrt{p_{b(k)}} \eta_{b(k)}(t)}{\sqrt{1 + p_{b(k)}}} + \frac{\varepsilon_k(t)}{\sqrt{1 + p_{b(k)}}}$$

B branches, sizes κ_b , $b = 1, \dots, B$
 κ companies in no branch

$$\rho_T^{(q)}(\lambda) = (K - \kappa - B)G \left(\lambda - \mu_B^{(q)}, \frac{(v_B^{(q)})^2}{T} \right) + \kappa G \left(\lambda - 1, \frac{(v_0^{(q)})^2}{T^q} \right) \\ + \sum_{b=1}^B \delta \left(\lambda - \left(1 + (\kappa_b - 1) \left(\frac{p_b}{1 + p_b} \right)^q \right) \right)$$

where

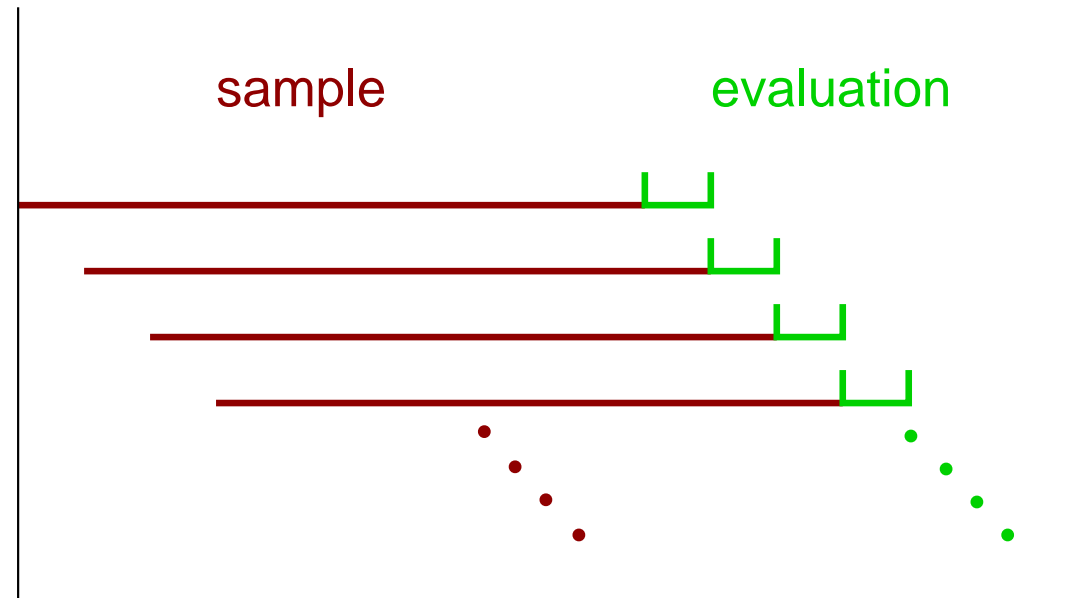
$$\mu_B^{(q)} = 1 - \frac{1}{B} \sum_{b=1}^B \left(\frac{p_b}{1 + p_b} \right)^q$$

Application to Market Data

Markowitz Optimization after Noise Reduction

portfolio optimization

Markowitz theory



Swedish stock returns

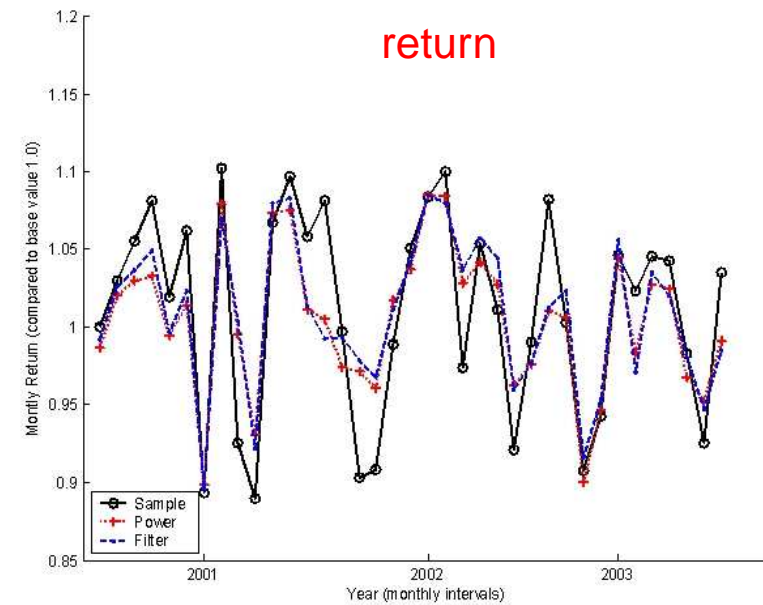
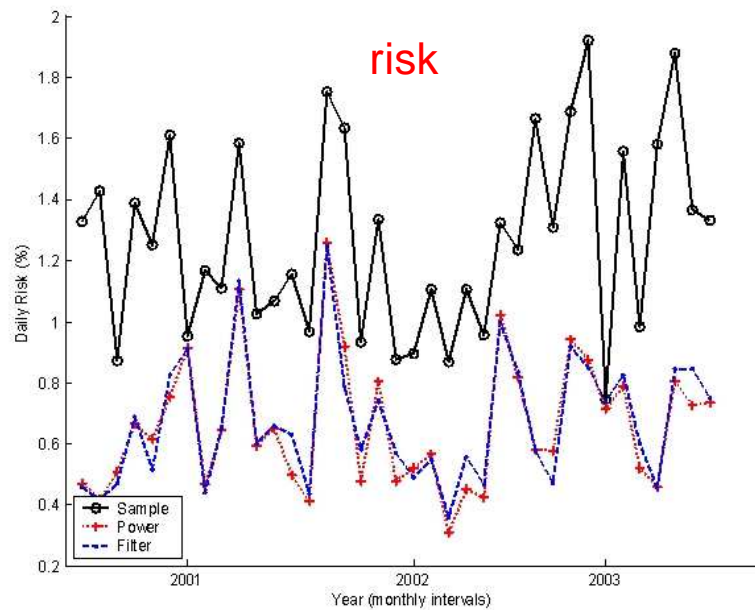
197 companies, daily, from July 12, 1999 to July 18, 2003
sample: one year — evaluation steps: one week

Standard & Poor's 500

100 most actively traded stocks, daily data 2002 to 2006
sample: 150 days — evaluation steps: 14 days

Swedish Stocks — No Constraints

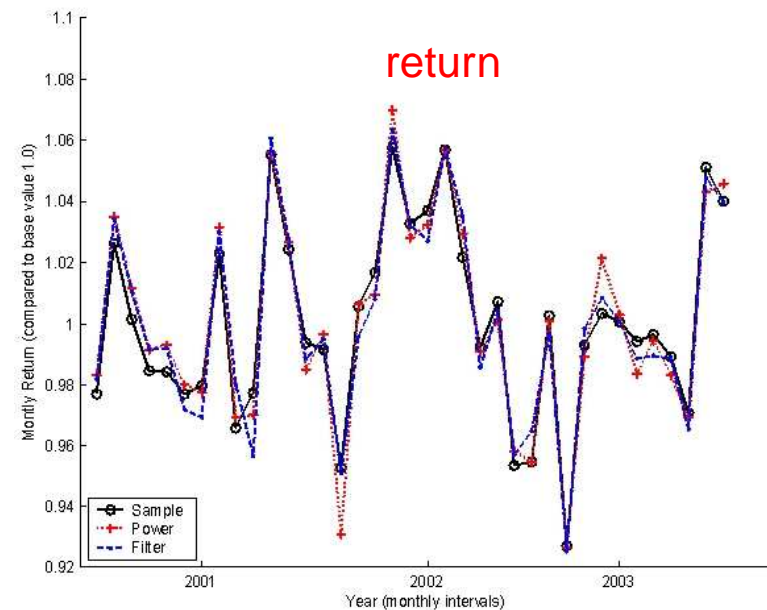
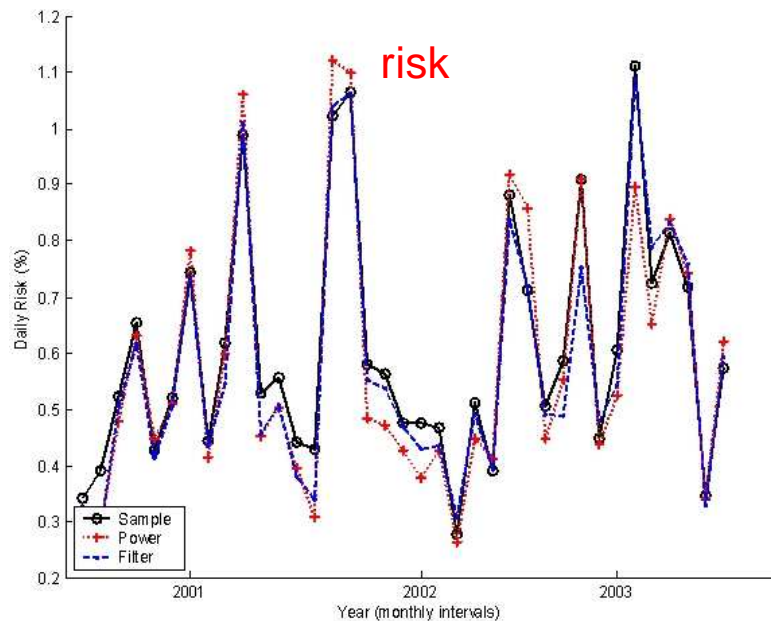
Markowitz theory with desired return of 0.3% per week



	yearly actual risk [%]	yearly actual return [%]
sample	20.7	11.1
power mapped	11.3	5.0
filtered	11.4	10.5

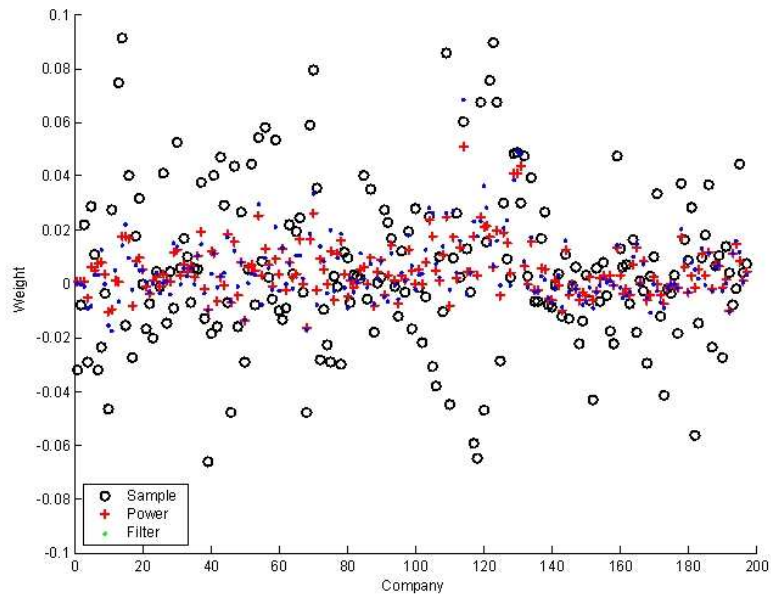
Swedish Stocks — Constraint: No Short Selling

Markowitz theory with desired return of 0.1% per week

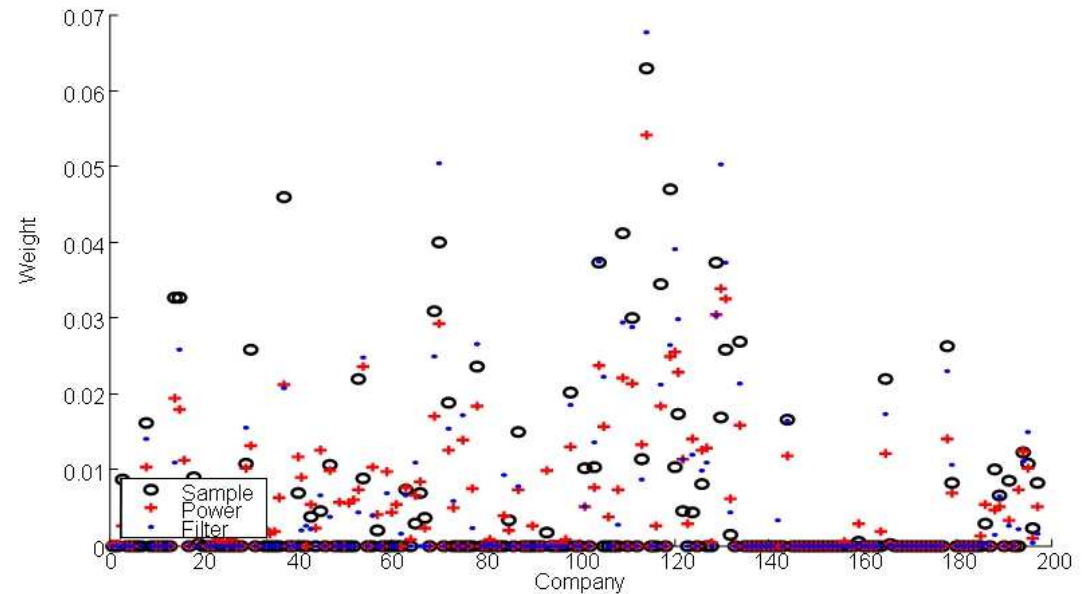


	yearly actual risk [%]	yearly actual return [%]
sample	10.1	0.5
power mapped	9.9	1.1
filtered	9.9	0.7

Swedish Stocks — Weights



no constraints



constraint: no short selling

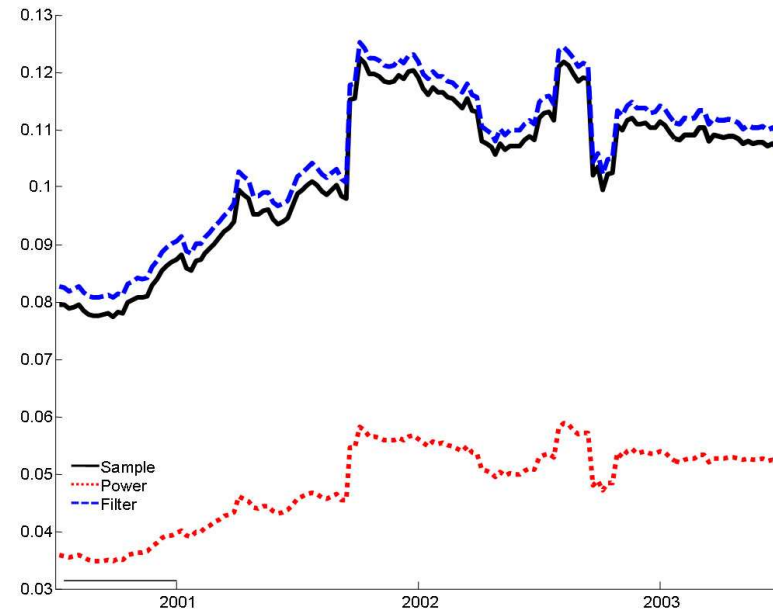
→ less rigid: filtering seems favored
more rigid: power mapping seems favored

Mean Value of Correlation Matrix

in a sampling period

$$c = \frac{1}{K^2} \sum_{k,l} C_{kl}$$

very similar curves!



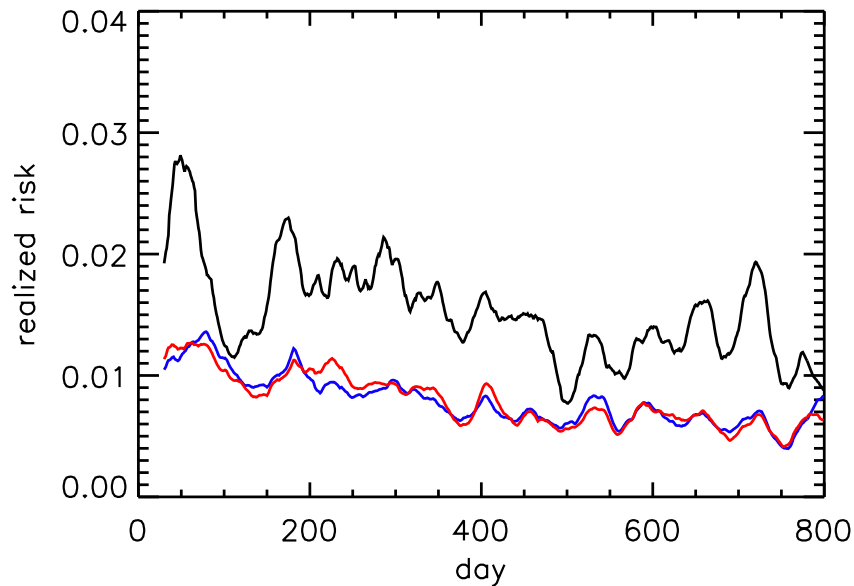
important: Markowitz optimization is invariant under scaling

$$C \longrightarrow \gamma C \text{ for all } \gamma > 0$$

power mapped C can be readjusted with $c^{(\text{original})} / c^{(\text{power mapped})}$

Standard & Poor's — Adjusted Power

$K = 100, T = 150, q_{\text{opt}} = 1.8$



constraint: no short selling

Mean realized risk

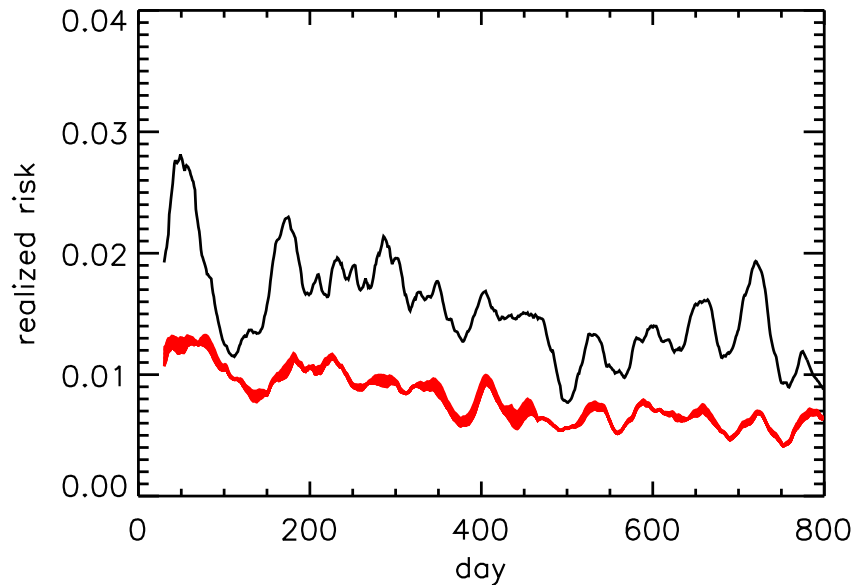
C^{sample}	1.548e-2
C^{filter}	0.809e-2
$C(q)$	0.812e-2

Mean realized return

C^{sample}	0.42e-4
C^{filter}	5.45e-4
$C(q)$	5.90e-4

Standard & Poor's — Varying Power

$$K = 100, T = 150, q_{\text{opt}} = 1.8$$



constraint:
no short selling

Power-mapping yields good risk-reduction for wide range of q values

Summary and Conclusions

- portfolio risk depends on correlations
- Markowitz optimization is an Euler–Lagrange problem
- correlations are noise dressed
- two noise reduction methods discussed: filtering and power mapping
- both are good at reducing risk, perform differently in the presence of constraints (no short selling)
- example for everyday work in financial industry